

INVESTIGATION OF QUEUEING NETWORK WITH UNRELIABLE SYSTEMS AND LARGE NUMBER OF MESSAGES

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Abstract. The Markov network with unreliable queueing systems and a large number of messages is investigated. The service channels of systems are exposed to random failure, besides the time of the proper functionality and the time of reconstruction of each channel of system has the exponential distribution with distinctive parameters. The system of difference-differential equations of Kolmogorov for the states probabilities is compiled. The partial differential equation for the probability density function of the vector of states is deduced. The systems of ordinary differential equations for an average number of messages and serviceable channels of network systems are received.

Introduction

Let us examine the closed exponential queueing network with the K messages of the same type which consist of $n+1$ queueing systems (QS) S_1, S_2, \dots, S_n . The system S_i includes m_i identical service channels, $i = \overline{1, n}$, and $m_0 = K$.

Considering that the service channels of the system S_0 are absolutely reliable and in the other systems S_1, S_2, \dots, S_n the service channels are exposed to random failure; besides the time of the proper functionality of each S_i system's channel has the exponential distribution with the parameter β_i , $i = \overline{1, n}$. After the breakage, the channel starts to reconstruct immediately. The time of reconstruction also has exponential distribution with the parameter γ_i , $i = \overline{1, n}$. After servicing in system S_i the message immediately transfers into the system S_j with probability p_{ij} ,

$i, j = \overline{0, n}$, $p_{00} = 0$, $\sum_{j=0}^n p_{ij} = 1$, $i = \overline{0, n}$. The matrix $P = \|p_{ij}\|_{(n+1) \times (n+1)}$ is transition probability matrix of irreducible Markov chains. If it arrives in the system S_j message finds at least one service channel operable and free from the other messages it is immediately serviced and the time of service is a random variable with

the parameter μ_i , $i = \overline{1, n}$. Otherwise the message expects the beginning of service without restriction on duration of waiting. Let's assume that if the service channel would fail while completing some message, then after the restoration the interrupted message will be completed. Disciplines of the message processing in the network systems are FIFO. Assuming that the service time of messages, durations of serviceable work of channels and restoration time of service channels are independent random variables.

Our aim is to receive the system of the differential equations for the average number of messages and serviceable channels in the network QS at the large values of K . It should be noted that the presented techniques of the results reception has been offered for the first time in the works [1, 2] for the exponential networks without the specified features (with reliable QS).

1. The system of equations for the states probabilities

The state of such a network at the moment t could be described through vector

$$z(t) = (d(t), k(t)) = (d_1(t), d_2(t), \dots, d_n(t), k_1(t), k_2(t), \dots, k_n(t)), \quad (1)$$

where $d_i(t)$ and $k_i(t)$ are the numbers of serviceable channels and the messages numbers in the system S_i at the moment t accordingly, $0 \leq d_i(t) \leq m_i$, $0 \leq k_i(t) \leq K$, $t \in [0, +\infty)$, $i = \overline{1, n}$. It is obvious that $k_0(t) = K - \sum_{i=1}^n k_i(t)$ – is the number of messages in the system S_0 at the moment t .

Vector $z(t)$ describes $2n$ - dimensional Markov process with the continuous time and the definite number of states. Let's consider, that

$$P(d, k, t) = P(d(t) = d, k(t) = k),$$

where $d = (d_1, d_2, \dots, d_n)$, $0 \leq d_i \leq m_i$ and $k = (k_1, k_2, \dots, k_n)$, $0 \leq k_i \leq K$, $i = \overline{1, n}$. Let's denote I_i as n - vector with zero components excluding i , that is equals to 1. Let's describe the possible passages of Markov process $z(t)$ in the state $z(t + \Delta t) = (d, k, t + \Delta t)$ at the time Δt :

- from the state (d, k, t) the passage is possible with probability

$$1 - \left[\mu_0 \left(K - \sum_{i=1}^n k_i(t) \right) + \sum_{i=1}^n \left[\mu_i \min(m_i, k_i(t)) + \beta_i d_i(t) + \gamma_i (m_i - d_i(t)) \right] \right] \Delta t + o(t);$$

- from the state $(d, k - I_i, t)$ with probability

$$\left[\mu_0 p_{0i} \left(K - \sum_{i=1}^n k_i(t) + 1 \right) \Delta t + o(\Delta t) \right] \left[1 - \left\{ \mu_0 \sum_{\substack{j=1 \\ j \neq i}}^n p_{0j} (K - k_j(t)) + \sum_{j=1}^n \left[\mu_j \min(d_j(t), k_j(t)) + \beta_j d_j(t) + \gamma_j (m_j - d_j(t)) \right] \right\} \Delta t + o(\Delta t) \right], \quad i = \overline{1, n};$$

- from the state $(d, k + I_i, t)$ with probability

$$\left[\mu_i p_{i0} \min(d_i(t), k_i(t) + 1) \Delta t + o(\Delta t) \right] \left[1 - \left\{ \mu_0 \left(K - \sum_{j=1}^n k_j(t) + 1 \right) + \sum_{\substack{j=1 \\ j \neq i}}^n \left[\mu_j \min(d_j(t), k_j(t)) + \beta_j d_j(t) + \gamma_j (m_j - d_j(t)) + \mu_i \min(d_i(t), k_i(t)) + \beta_i d_i(t) + \gamma_i (m_i - d_i(t)) \right] \right\} \Delta t + o(\Delta t) \right], \quad i = \overline{1, n};$$

- from the state $(d, k + I_i - I_j, t)$ with probability

$$\left[\mu_i p_{ij} \min(d_i(t), k_i(t) + 1) \Delta t + o(\Delta t) \right] \left[1 - \left\{ \mu_0 \left(K - \sum_{j=1}^n k_j(t) \right) + \sum_{\substack{r=1 \\ r \neq i, j}}^n \left[\mu_r \min(d_r(t), k_r(t)) + \beta_r d_r(t) + \gamma_r (m_r - d_r(t)) \right] + \mu_i \min(d_i(t), k_i(t)) + \beta_i d_i(t) + \gamma_i (m_i - d_i(t)) + \mu_j \min(d_j(t), k_j(t) + 1) + \beta_j d_j(t) + \gamma_j (m_j - d_j(t)) \right\} \Delta t + o(\Delta t) \right], \quad i, j = \overline{1, n};$$

- from the state $(d - I_i, k, t)$ with probability

$$\left[\gamma_i (m_i - d_i(t) + 1) \Delta t + o(\Delta t) \right] \left[1 - \left\{ \mu_0 \left(K - \sum_{j=1}^n k_j(t) \right) + \right.$$

$$+ \sum_{j=1}^n \left[\mu_j \min(d_j(t), k_j(t)) + \beta_j d_j(t) + \gamma_j (m_j - d_j(t)) \right] \Delta t + o(\Delta t) \Big\}, \quad i = \overline{1, n};$$

- from the state $(d + I_i, k, t)$ with probability

$$\begin{aligned} & \left[\beta_i (d_i(t) + 1) \Delta t + o(\Delta t) \right] \left[1 - \left\{ \mu_0 \left(K - \sum_{j=1}^n k_j(t) \right) + \right. \right. \\ & + \sum_{\substack{j=1 \\ j \neq i}}^n \left[\mu_j \min(d_j(t), k_j(t)) + \beta_j d_j(t) + \gamma_j (m_j - d_j(t)) \right] + \mu_i \min(d_i(t) + 1, k_i(t)) + \\ & \left. \left. + \gamma_i (m_i - (d_i(t) + 1)) \right\} \Delta t + o(\Delta t) \right] \Big\}, \quad i = \overline{1, n}; \end{aligned}$$

- from all other states with probability $o(\Delta t)$.

Then, the usage of the formula of total probability makes it possible to write the system of difference equations for the states probabilities

$$\begin{aligned} P(d, k, t + \Delta t) = & \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} \min(d_i(t), k_i(t) + 1) P(d, k + I_i - I_j, t) \Delta t + \\ & + \mu_0 \left(K - \sum_{i=1}^n k_i(t) + 1 \right) P(d, k - I_j, t) \Delta t + \\ & + \sum_{i=1}^n \mu_i p_{i0} \min(d_i(t), k_i(t) + 1) P(d, k + I_i, t) \Delta t + \\ & + \sum_{j=1}^n \gamma_i (m_i - d_i(t) + 1) P(d - I_j, k, t) \Delta t + \sum_{i=1}^n \beta_i (d_i(t) + 1) P(d + I_i, k, t) \Delta t + \\ & + \left\{ 1 - \left[\mu_0 \left(K - \sum_{i=1}^n k_i(t) \right) + \sum_{i=1}^n \mu_i p_{i0} \min(d_i(t), k_i(t)) + \sum_{i=1}^n \gamma_i (m_i - d_i(t)) + \right. \right. \\ & \left. \left. + \sum_{i=1}^n \beta_i d_i(t) \right] \Delta t \right\} P(d, k, t) + o(\Delta t), \end{aligned}$$

from which at $\Delta t \rightarrow 0$ we receive the system of difference-differential equations of Kolmogorov for the states probabilities

$$\begin{aligned}
\frac{dP(d, k, t)}{dt} = & \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} \min(d_i(t), k_i) [P(d, k + I_i - I_j, t) - P(d, k, t)] + \\
& + \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} [\min(d_i(t), k_i(t) + 1) - \min(d_i(t), k_i(t))] P(d, k + I_i - I_j, t) + \\
& + \mu_0 \left(K - \sum_{i=1}^n k_i(t) \right) [P(d, k - I_j, t) - P(d, k, t)] + \mu_0 P(d, k - I_j, t) + \\
& + \sum_{i=1}^n \mu_i p_{i0} \min(d_i(t), k_i(t)) [P(d, k + I_i, t) - P(d, k, t)] + \\
& + \sum_{i=1}^n \mu_i p_{i0} [\min(d_i(t), k_i(t) + 1) - \min(d_i(t), k_i(t))] P(d, k + I_i, t) + \\
& + \sum_{j=1}^n \gamma_j (m_j - d_j(t)) [P(d - I_j, k, t) - P(d, k, t)] + \sum_{i=1}^n \gamma_i P(d - I_i, k, t) + \\
& + \sum_{i=1}^n \beta_i d_i(t) [P(d + I_i, k, t) - P(d, k, t)] + \sum_{i=1}^n \beta_i P(d + I_i, k, t). \quad (2)
\end{aligned}$$

The solution of this system in the analytical form is generally inconvenient. Therefore we will consider the important case of the large number of messages in the network, $K \gg 1$. In order to determine probability distribution of the random vector $z(t)$, it is convenient to switch to the relative variables, considering vector $\xi(t) = \left(\frac{d_1(t)}{K}, \frac{d_2(t)}{K}, \dots, \frac{d_n(t)}{K}, \frac{k_1(t)}{K}, \frac{k_2(t)}{K}, \dots, \frac{k_n(t)}{K} \right)$. In this case possible values of this vector at the fixed t will belong to the bounded closed set

$$G = \left\{ (y, x) = (y_1, y_2, \dots, y_n, x_1, x_2, \dots, x_n) : x_i \geq 0, \sum_{i=1}^n x_i \leq 1, 0 \leq y_i \leq \frac{m_i}{K} \right\}, \quad (3)$$

in which they place in the nodes of the $2n$ -dimensional grid at the distance $\varepsilon = \frac{1}{K}$ from each other. While magnifying K "the charging density" of set G with the possible components of vector $\xi(t)$ will increase, and it is possible to consider, that it has a continuous distribution with the probabilities density $p(y, x, t)$, and $K^{2n} P(d, k, t) \xrightarrow{K \rightarrow \infty} p(y, x, t)$. Therefore it is possible to use the approximation of the function $P(d, k, t)$, using the relation $K^{2n} P(d, k, t) = K^{2n} P(yK, xK, t) = p(y, x, t)$, $(y, x) \in G$.

Let denote that $e_i = \varepsilon I_i$, $i = \overline{1, n}$, $c(b) = \begin{cases} 1, & b > 0 \\ 0, & b \leq 0 \end{cases}$, and

$$\min(b, a+1) = \min(b, a) + c(b-a), \quad c(b-a) = \frac{\partial \min(b, a)}{\partial a}, \quad (4)$$

thus $\min(b, a) = \begin{cases} a, & b \geq a \\ b, & b < a \end{cases}$. Using the relative variables $y_i = \frac{d_i}{K}$, $x_i = \frac{k_i}{K}$, $l_i = \frac{m_i}{K}$, $i = \overline{1, n}$, expression (4) and that if $K \rightarrow \infty$, $\varepsilon \rightarrow 0$, system (2) can be written as follows:

$$\begin{aligned} \frac{\partial p(y, x, t)}{\partial t} = & \sum_{i=1}^n \sum_{j=1}^n K \mu_i p_{ij} \min(y_i, x_i) [p(y, x + e_i - e_j, t) - p(y, x, t)] + \\ & + \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} \frac{\partial \min(y_i, x_i)}{\partial x_i} p(y, x + e_i - e_j, t) + \\ & + K \mu_0 \left(1 - \sum_{i=1}^n x_i \right) [p(y, x - e_j, t) - p(y, x, t)] + \mu_0 (py, x - e_j, t) + \\ & + \sum_{i=1}^n K \mu_i p_{i0} \min(y_i, x_i) [p(y, x + e_i, t) - (py, x, t)] + \\ & + \sum_{i=1}^n \mu_i p_{i0} \frac{\partial \min(y_i, x_i)}{\partial x_i} p(y, x + e_i, t) + \\ & + \sum_{i=1}^n K \gamma_i (l_i - y_i) [p(y - e_i, x, t) - p(y, x, t)] + \sum_{j=1}^n \gamma_j p(y - e_j, x, t) + \\ & + \sum_{i=1}^n K \beta_i y_i [p(y + e_i, x, t) - p(y, x, t)] + \sum_{i=1}^n \beta_i p(y + e_i, x, t). \end{aligned} \quad (5)$$

2. The system of differential equations for expected characteristics

Let's present the right part (5) with the accuracy of term ε^2 . If $p(y, x, t)$ is twice continuously differentiated at y and x , than

$$\begin{aligned} p(y, x \pm e_i, t) = & p(y, x, t) \pm \varepsilon \frac{\partial p(y, x, t)}{\partial x_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial x_i^2} + o(\varepsilon^2), \\ p(y, x + e_i - e_j, t) = & p(y, x, t) + \varepsilon \left(\frac{\partial p(y, x, t)}{\partial x_i} - \frac{\partial p(y, x, t)}{\partial x_j} \right) + \end{aligned}$$

$$\begin{aligned}
& + \frac{\varepsilon^2}{2} \left(\frac{\partial^2 p(y, x, t)}{\partial x_i^2} - 2 \frac{\partial^2 p(y, x, t)}{\partial x_i \partial x_j} + \frac{\partial^2 p(y, x, t)}{\partial x_j^2} \right) + o(\varepsilon^2), \\
p(y \pm e_i, x, t) & = p(y, x, t) \pm \varepsilon \frac{\partial p(y, x, t)}{\partial y_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial y_i^2} + o(\varepsilon^2), \quad i = \overline{1, n}. \quad (6)
\end{aligned}$$

Using them and that $\varepsilon K = 1$, we receive

$$\begin{aligned}
\frac{\partial p(y, x, t)}{\partial t} & = \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} \min(y_i, x_i) \left[\left(\frac{\partial p(y, x, t)}{\partial x_i} - \frac{\partial p(y, x, t)}{\partial x_j} \right) + \right. \\
& \quad \left. + \frac{\varepsilon}{2} \left(\frac{\partial^2 p(y, x, t)}{\partial x_i^2} - 2 \frac{\partial^2 p(y, x, t)}{\partial x_i \partial x_j} + \frac{\partial^2 p(y, x, t)}{\partial x_j^2} \right) \right] + \\
& + \sum_{i=1}^n \sum_{j=1}^n \mu_i p_{ij} \frac{\partial \min(y_i, x_i)}{\partial x_i} \left[p(y, x, t) + \varepsilon \left(\frac{\partial p(y, x, t)}{\partial x_i} - \frac{\partial p(y, x, t)}{\partial x_j} \right) + \right. \\
& \quad \left. + \frac{\varepsilon^2}{2} \left(\frac{\partial^2 p(y, x, t)}{\partial x_i^2} - 2 \frac{\partial^2 p(y, x, t)}{\partial x_i \partial x_j} + \frac{\partial^2 p(y, x, t)}{\partial x_j^2} \right) \right] + \\
& + \sum_{i=1}^n \mu_0 p_{0i} \left(1 - \sum_{i=1}^n x_i \right) \left[- \frac{\partial p(y, x, t)}{\partial x_i} + \frac{\varepsilon}{2} \frac{\partial^2 p(y, x, t)}{\partial x_i^2} \right] + \\
& + \mu_0 \left[p(y, x, t) - \varepsilon \frac{\partial p(y, x, t)}{\partial x_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial x_i^2} \right] + \\
& + \sum_{i=1}^n \mu_i p_{i0} \min(y_i, x_i) \left[\frac{\partial p(y, x, t)}{\partial x_i} + \frac{\varepsilon}{2} \frac{\partial^2 p(y, x, t)}{\partial x_i^2} \right] + \\
& + \sum_{i=1}^n \mu_i p_{i0} \frac{\partial \min(y_i, x_i)}{\partial x_i} \left[p(y, x, t) + \varepsilon \frac{\partial p(y, x, t)}{\partial x_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial x_i^2} \right] + \\
& + \sum_{i=1}^n \gamma_i (l_i - y_i) \left[- \frac{\partial p(y, x, t)}{\partial y_i} + \frac{\varepsilon}{2} \frac{\partial^2 p(y, x, t)}{\partial y_i^2} \right] + \\
& + \sum_{i=1}^n \gamma_i \left[p(y, x, t) - \varepsilon \frac{\partial p(y, x, t)}{\partial y_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial y_i^2} \right] + \\
& + \sum_{i=1}^n \beta_i y_i \left[\frac{\partial p(y, x, t)}{\partial y_i} + \frac{\varepsilon}{2} \frac{\partial^2 p(y, x, t)}{\partial y_i^2} \right] +
\end{aligned}$$

$$+ \sum_{i=1}^n \beta_i \left[p(y, x, t) + \varepsilon \frac{\partial p(y, x, t)}{\partial y_i} + \frac{\varepsilon^2}{2} \frac{\partial^2 p(y, x, t)}{\partial y_i^2} \right] + o(\varepsilon^2).$$

Thus, the density $p(y, x, t)$ is satisfied with the accuracy within the term ε^2 to the Kolmogorov-Fokker-Planc equation:

$$\begin{aligned} \frac{\partial p(y, x, t)}{\partial t} = & - \sum_{i=1}^n \frac{\partial}{\partial y_i} (A_i^{(1)}(y) p(y, x, t)) - \sum_{i=1}^n \frac{\partial}{\partial x_i} (A_i^{(2)}(y, x) p(y, x, t)) + \\ & + \frac{\varepsilon}{2} \sum_{i,j=1}^n \frac{\partial^2}{\partial y_i \partial y_j} (B_{ij}^{(1)}(y) p(y, x, t)) + \frac{\varepsilon}{2} \sum_{i,j=1}^n \frac{\partial^2}{\partial x_i \partial x_j} (B_{ij}^{(2)}(y, x) p(y, x, t)), \end{aligned} \quad (7)$$

where

$$A_i^{(1)}(y) = \gamma_i (l_i - y_i) - \beta_i y_i, \quad (8)$$

$$A_i^{(2)}(y, x) = \sum_{j=1}^n \mu_j p_{ji}^* \min(y_j, x_j) + \mu_0 p_{0i} \left(1 - \sum_{i=1}^n x_i \right), \quad i = \overline{1, n}, \quad (9)$$

$$p_{ji}^* = \begin{cases} p_{ji}, & j \neq i, \\ p_{ii} - 1, & j = i; \end{cases} \quad B_{ii}^{(1)}(y) = \gamma_i (l_i - y_i) + \beta_i y_i; \quad B_{ij}^{(1)}(y) = 0, \quad i \neq j;$$

$$B_{ii}^{(2)}(y, x) = \sum_{j=1}^n \mu_j p_{ji}^{**} \min(y_j, x_j) + \mu_0 p_{0i} \left(1 - \sum_{i=1}^n x_i \right),$$

$$p_{ji}^{**} = \begin{cases} p_{ji}, & j \neq i, \\ 1 - p_{ii}, & j = i; \end{cases} \quad B_{ij}^{(2)}(y, x) = -\mu_i p_{ij} \min(y_i, x_i), \quad i \neq j, \quad i = \overline{1, n}.$$

As the density $p(y, x, t)$ satisfies the Kolmogorov-Fokker-Planc equation and $A_i^{(1)}(y)$, $A_i^{(2)}(y, x)$ piecewise linear functions on y , x , according to [3], the mathematical expectations $w_i(t) = M \left\{ \frac{d_i(t)}{K} \right\}$, $n_i(t) = M \left\{ \frac{k_i(t)}{K} \right\}$, $i = \overline{1, n}$, with the accuracy within the terms of infinitesimal order $O(\varepsilon^2)$ are defined from the systems of the equations

$$\frac{dw_i(t)}{dt} = A_i^{(1)}(w(t)) = \gamma_i (l_i - w_i(t)) - \beta_i w_i(t), \quad i = \overline{1, n}, \quad (10)$$

$$\frac{dn_i(t)}{dt} = A_i^{(2)}(w(t), n(t)) = \sum_{j=1}^n \mu_j p_{ji}^* \min(w_j(t), n_j(t)) + \mu_0 p_{0i} \left(1 - \sum_{i=1}^n n_i(t)\right), \quad (11)$$

The right hand sides of system (11) are continuous piecewise linear functions. By segmentation of phase space and obtaining solutions of system (11) in ranges of right hand sides linearity it is possible to solve the whole system.

Let $\Omega(t) = \{1, 2, \dots, n\}$ - be set of vector $n(t)$ component indices. Let's divide $\Omega(t)$ into two disjoint sets $\Omega_0(t)$ and $\Omega_1(t)$:

$$\Omega_0(t) = \{i : w_i(t) < n_i(t) \leq 1\}, \quad \Omega_1(t) = \{j : 0 \leq n_j(t) \leq w_j(t)\}.$$

Each partitioning specifies disjoint regions $G_\tau(t)$ in the set

$$G(t) = \left\{ n(t) : n_i(t) \geq 0, \sum_{i=1}^n n_i(t) \leq 1 \right\}$$

such that:

$$G_\tau(t) = \left\{ n(t) : w_i(t) < n_i(t) \leq 1, i \in \Omega_0(t); \right. \\ \left. 0 \leq n_j(t) \leq w_j(t), j \in \Omega_1(t); \sum_{c=1}^n n_c(t) \leq 1 \right\}, \quad \tau = 1, 2, \dots, 2^n, \quad \bigcup_{\tau=1}^{2^n} G_\tau(t) = G(t).$$

Then system of equations (11) of explicit form for each region $G_\tau(t)$ is:

$$\frac{dn_i(t)}{dt} = \sum_0 \mu_j p_{ji}^* w_j(t) + \sum_1 \mu_j p_{ji}^* n_j(t) + \mu_0 p_{0i} \left(1 - \sum_{i=1}^n n_i(t)\right), \quad i = \overline{1, n}, \quad (12)$$

where $\sum_0 = \sum_{j \in \Omega_0(t)}$, $\sum_1 = \sum_{j \in \Omega_1(t)}$. The solution of the system of equations (10), (12) makes it possible to obtain an average relative number of messages and serviceable channels at any QS of queuing network. At the region A : $\Omega_0(t) = \{\emptyset\}$, $\Omega_1(t) = \{1, 2, \dots, n\}$, when queues in average at the system are absent, systems (10), (12) look like

$$\frac{dw_i(t)}{dt} = \gamma_i (l_i - w_i(t)) - \beta_i w_i(t) = \gamma_i l_i - (\gamma_i + \beta_i) w_i(t), \quad (13)$$

$$\frac{dn_i(t)}{dt} = \sum_1 \mu_j p_{ji}^* n_j(t) + \mu_0 p_{0i} \left(1 - \sum_{i=1}^n n_i(t)\right). \quad (14)$$

3. Example

Let's consider the closed exponential queueing network with unreliable service channels and central QS, $n = 4$, $K = 1000$; $m_1 = 11$, $m_2 = 9$, $m_3 = 10$, $m_4 = 15$. The intensity of arrival of messages from outside medium (system S_0) is $\mu_0 = 5$. Intensity of service of messages in service channels of QS are $\mu_1 = 1.5$, $\mu_2 = 1$, $\mu_3 = 1.4$, $\mu_4 = 0.9$. Averages of duration of serviceable work for each service channel of QS $\beta_1^{-1} = 1.65$, $\beta_2^{-1} = 1.41$, $\beta_3^{-1} = 1.4$, $\beta_4^{-1} = 1.31$ are equal accordingly. Averages of duration of reconstruction of faulty service channels of QS are $\gamma_1^{-1} = 1.31$, $\gamma_2^{-1} = 0.9$, $\gamma_3^{-1} = 1.2$, $\gamma_4^{-1} = 1.65$. Probabilities of transition of messages: $p_{01} = p_{02} = p_{03} = p_{04} = p_{10} = p_{20} = p_{30} = p_{40} = 1/4$, $p_{14} = p_{24} = p_{34} = 3/4$, $p_{14} = 1/2$, $p_{42} = p_{43} = 1/8$, other $p_{ij} = 0$, $i, j = \overline{0,4}$. At the initial moment of time $t = 0$: $n_i(0) = 0$, $w_i(0) = m_i$, and $\min(w_i(t), n_i(t)) = n_i(t)$, $i = \overline{1,4}$. Then the system (14) is a system of non-homogeneous linear differential equations and can be written in the vector form

$$\frac{dn(t)}{dt} = An(t) + Q(t), \quad (15)$$

where $n^T(t) = (n_1(t), n_2(t), \dots, n_5(t))$ - vector of average relative number of messages in each of QS. The decision of system (14) can be found as

$$n(t) = e^{At}n(0) + \int_0^t e^{A(t-\tau)}Q(\tau)d\tau.$$

Analytical expressions for change of average relative number of messages in network systems look like

$$\begin{aligned} n_1(t) &= 0.002 + 0.0031e^{-1.9998t} - 0.003e^{-5.3195t} - 0.0017e^{-1.4196t} + 0.0005e^{-1.0609t}, \\ n_2(t) &= 0.0012 + 0.0002e^{-1.9998t} - 0.0027e^{-5.3195t} + 0.0001e^{-1.4196t} + 0.0011e^{-1.0609t}, \\ n_3(t) &= 0.0009 + 0.0004e^{-1.9998t} - 0.0003e^{-5.3195t} + 0.0019e^{-1.4196t} + 0.0002e^{-1.0609t}, \\ n_4(t) &= 0.0053 - 0.0039e^{-1.9998t} - 0.0007e^{-5.3195t} - 0.0003e^{-1.4196t} - 0.0004e^{-1.0609t}. \end{aligned}$$

When these expressions are multiplied on K we will receive expressions for an average of messages in QS $N_i(t) = Kn_i(t)$, $i = \overline{1,4}$. Figures of change $N_1(t)$, $N_3(t)$ are presented in Figures 1 and 2.

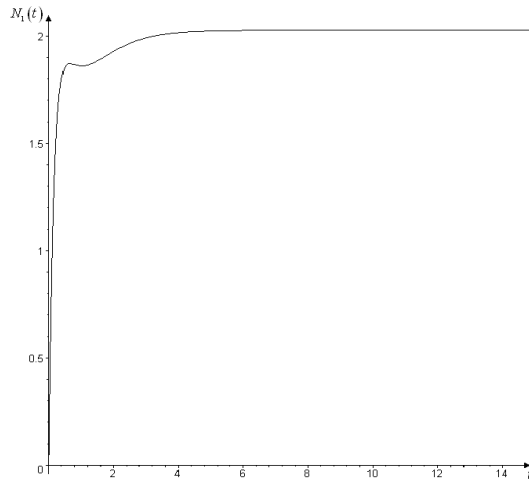


Fig. 1. Change of average number of messages in QS S_1

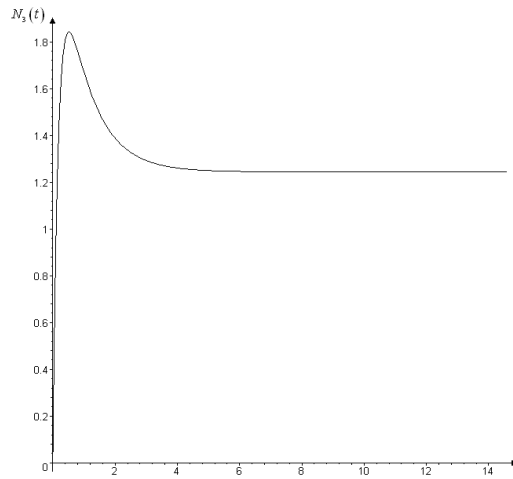


Fig. 2. Change of average number of messages in QS S_3

Solving system (13) we will receive expressions for the average relative number of serviceable channels in network systems:

$$w_1(t) = 0.0061 + 0.0049e^{-1.3694t}, w_2(t) = 0.0055 + 0.0035e^{-1.8203t},$$

$$w_3(t) = 0.0054 + 0.0046e^{-1.5476t}, w_4(t) = 0.0066 + 0.0084e^{-1.3694t}.$$

from $d_i(t) = Kw_i(t)$, $i = \overline{1,4}$. Figures of change $d_1(t)$, $d_3(t)$ are presented in Figures 3 and 4.

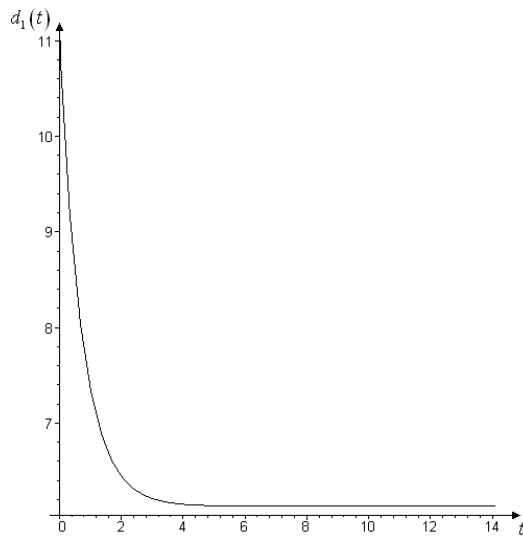


Fig. 3. Change of average number of serviceable channels in QS S_1

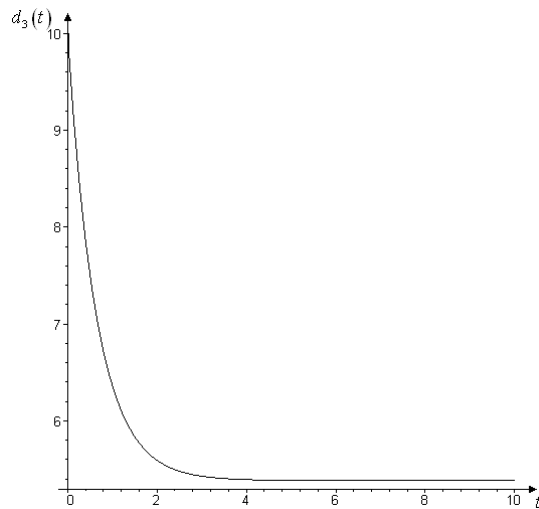


Fig. 4. Change of average number of serviceable channels in QS S_3

Conclusions

The considered method of diffusive approximations for finding the average number of messages and serviceable channels of QN makes it possible to precisely define of the given average characteristics in the stationary and transient regime. The accuracy of method increases with an increase in a number of messages in the network.

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